Correlated Data 4/12/22 - Paired Oak - Binan lespose STAT 216 - Paired t-lest: - Quantitative (normal) vespise - Pre/Post - Husbard/With - Measured in pairs y = go for welks D D 1 2×/week nours ~ cols (See. 8.1) McNemar's Test of independence for correlated Proportions: $H_s: P(Y_1 = Y_{es}) = P(Y_2 = Y_{es})$ Ha: PLY, = Yes) P(Yz = Yes) Dre of < , > , 7 dependents on research queel. This = P(being in the in 1000 - in col) $\Rightarrow P(Y_{1} = Y_{es}) = \pi_{11} + \pi_{12} = \pi_{1+} + P(Y_{2} = Y_{es}) = \pi_{11} + \pi_{21} = \pi_{+1} + \pi_{21} = \pi_{+1} + \pi_{22} = \pi_{+1} + \pi_{-1} + \pi_{-1} = \pi_{+1} + \pi_{-1} + \pi_{-1} = \pi_{+1} + \pi_{$

- Only interested in "discardont" pairs -individuals in (1,2)th cell or (2,1)th cell We say there is marginal homogeneity if $\pi_{i2} = \pi_{2i}$ $Uhy? TT_{1+} = TT_{+1} \longrightarrow$ $T_{1+} - T_{1+1} = (T_{11} + T_{12}) - (T_{11} + T_{21}) = T_{12} - T_{21}$ Let Nt = N12 + N21 -> total # of discardon-pairs Then Under Ho, Niz ~ Bin(nt, 0.5). - Use to get puedue or no >10 (artsimer, me of think), $Z = N_{12} - N^{*}(0.5) = N_{12} - N_{21} \sim N(q_1)$ $\int \Lambda^{*}(0,s)(1-0,s) = \int \Pi_{12} + \Pi_{21}$ $z^2 \sim \chi^2(0) \sim R$ sr Example: (Pediatrics, 2006) - Fire Safety Researchess compared children's reactions to two kinds of alarms. () Conventional Smoke alarm (2) Recording of mathen's voice Sayins the child's name & orging then to wake up -> each divid tested by each alorn -> 2 measurements per child

Matthere Usice Awake Did Nor Awake Convertional и <mark>(</mark>) и Awoke Mam 9) Jest 1 10 Old not awake 23 1 24 Ho: TIZ = TZI (P(Awake | Alcom 1) - P(Awake) T_{1+} $A(am_2)$ Ha: TI12 7 TT21 p-value = P(X=0 or X=9 | X~Bin(9,0.5)) Let X = N12 (prior to data callection) P: Sun (dbinon (c(0,9), 9,0.5) = 0.0039063=) Strong evidence that the probability of working differs between the two alarm types. Approximate pueles: $Z = N_{12} - N_{21} = 0 - 9 = -3$ Miztazi Ja P: 2 = pnom(-3) = 0.0026998or polying $((-3)^2, 1, (over, tri) = FALSE)$

Mother Usice Awake Did Not Awake Convertional Awoke 14 0 14 Alam 36 Oid not 9 test 10 avake l 23 24 How would the dara be organized ma 2x2 table if then were incorrectly treated as independent measurements. 1D Alam Type Awoke or Non Yes Conv.) Morrer Z 2 Awoke or Nor les No Conn Alam lype Mother 48

Asymptotic CI for TT1+ - TT1 : (dependent data) $SE(\pi_{12}-\pi_{11}) = \frac{1}{n} \left((n_{12}+n_{21}) - (n_{12}-n_{21})^2 \right)$ $CI = -\pi_{1+} - \pi_{+1} + \frac{1}{2^{k}} - 8\epsilon(\pi_{1+} - \pi_{+1})$ from N(o,) Asthna Example - 95% CT (0.01891, 0.03709) We are 95% confident that the the difference in propurning of 13-year-olds having asthma - zo-year-olds 'having astma (13-20) is between (0.019, 0.037), among children similar to those in the sample. We are 95% confident that the Tisk of asthra at age 13 is between 0.019 × 0.037 higher than the Tisk of asthra at age 20

Notes on "correct = FALSE" argument in R

- When using a normal approximation (continuers) on discrete data, we often employ a "continuity correction" -

 $Z_{c}^{2} = (1n_{12}-n_{21}-0.5)^{2}$ Mc Nenar : Pales $\tilde{\chi}_{1}^{2}$ $\tilde{\chi}_{1}^{2}$ $\overline{X}_{c}^{2} = \sum \left(\frac{\left(\text{obs-exp} \right) - 0.5 \right)^{2}}{e_{xp}}$ Chi-squared

More Generally - Modeling Framework Option: Population-averaged models - Marginal models $P(Y_1 = i) = \alpha + \delta$ $P(Y_t=1) = x + Sx_t$ $P(Y_2=1) = \alpha$ $S = P(Y_1 = 1) - P(Y_2 = 1)$ or another common model: $\log_i \left(P(Y_{t=1}) \right) = \log \left(\frac{P(Y_{t=1})}{1 - P(Y_{t=1})} \right) = \alpha + \delta X_t$ - But data are not independent. "Clustered data. - Need definer method for fitting these models -> generalized estimating equations (GEE) ---- "Naive SES" --- treating data as independent * "Robust SES" --- account for correlation

